

Recap:

$$\bar{X} = \begin{bmatrix} x_1 \\ \vdots \\ x_n \end{bmatrix} \sim f_{x_1, x_2, \dots, x_n}(x_1, \dots, x_n) \quad \bar{\mu} \quad K_{\bar{X}}$$

Assuming \bar{X} is zero-mean:

$$K_{\bar{X}} = Q \Lambda Q^T$$

$Q \in \mathbb{R}^{n \times n}$, orthonormal matrix, $Q^T Q = I$
 $\Lambda \in \mathbb{R}^{n \times n}$, diagonal.

$$\Lambda = \begin{bmatrix} \lambda_1 & & 0 \\ & \lambda_2 & \\ 0 & & \ddots \\ & & & \lambda_n \end{bmatrix}$$

$$\lambda_1 \geq \lambda_2 \geq \dots \geq \lambda_n$$

$$\bar{y} = Q^T \bar{X}$$

$$Q \bar{y} = Q Q^T \bar{X}$$

$$\bar{X} = Q \bar{y} = y_1 \bar{q}_1 + y_2 \bar{q}_2 + \dots + y_n \bar{q}_n$$

$$\bar{X}_{\text{app}} = \sum_{i=1}^m y_i \bar{q}_i$$

$$E[\|\bar{e}\|^2] = \sum_{i=m+1}^n \text{Var}(y_i)$$

$$= \sum_{i=m+1}^n \lambda_i$$

mean squared error.

$$Q = \begin{bmatrix} | & & | \\ \bar{q}_1 & & \bar{q}_n \\ | & & | \end{bmatrix}$$

If \bar{X} is not zero mean, \bar{X} has mean $\bar{\mu}$ and $K_{\bar{X}}$.

$\bar{X}_c = \bar{X} - \bar{\mu}$ has mean 0 and $K_{\bar{X}}$.

$$\bar{y} = Q^T \bar{X}_c$$

$$\bar{X}_c = Q \bar{y} = y_1 \bar{q}_1 + y_2 \bar{q}_2 + \dots + y_n \bar{q}_n$$

$$\bar{X}_c = \bar{\mu} + y_1 \bar{q}_1 + y_2 \bar{q}_2 + \dots + y_n \bar{q}_n$$

$$\bar{X}_{\text{app}} = \bar{\mu} + \sum_{i=1}^m y_i \bar{q}_i$$

Principal Component Analysis:

Data: $\bar{X}^{(1)}, \bar{X}^{(2)}, \bar{X}^{(3)}, \dots, \bar{X}^{(5000)}$

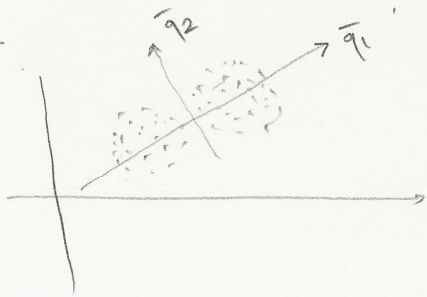
$\bar{X}^{(i)} \in \mathbb{R}^{1000}$

e.g. $\bar{X}^{(i)}$: image of human faces

$\bar{X}^{(i)}$: transcription of 1000 genes in cell i .

Data visualization.

$n=2$:



Recipe:

1) Estimate the covariance matrix of the data and the its mean $\hat{\mu} = \frac{1}{N} \sum_{i=1}^N \bar{x}_i$ $\hat{K}_{\bar{x}}$

2) Compute the eigendecomposition of $K_{\bar{x}}$
 $\hat{K}_{\bar{x}} = Q \Lambda Q^T$ $\bar{y} = Q(\bar{x} - \hat{\mu})$ $y_i = \bar{q}_i^T (\bar{x} - \hat{\mu})$
 $\bar{x}^{(i)} = y_1^{(i)} \bar{q}_1 + y_2^{(i)} \bar{q}_2 + \dots + y_n^{(i)} \bar{q}_n + \hat{\mu}$

\bar{q}_1 : first principal component.
 (the eigenvector corresponding to the largest eigenvalue λ_1)

the direction of maximal variation

\bar{q}_2 : second principal component.
 the eigenvector corresponding to λ_2

Properties of $K_{\bar{x}}$.

1) $K_{\bar{x}}$ is real $\left\{ \begin{array}{l} K_{\bar{x}} = Q \Lambda Q^T \\ \Lambda = \begin{bmatrix} \lambda_1 & & 0 \\ & \ddots & \\ 0 & & \lambda_n \end{bmatrix} \end{array} \right.$

2) $K_{\bar{x}}$ is symmetric

3) $\lambda_1 \geq \lambda_2 \geq \dots \lambda_n \geq 0 \Leftrightarrow K_{\bar{x}}$ is positive semi-definite

$\Leftrightarrow \bar{a}^T K_{\bar{x}} \bar{a} \geq 0$ for every $\bar{a} \in \mathbb{R}^n$.

\downarrow
 $\in \mathbb{R}^{n \times n}$

$$\bar{a}^T \mathbb{E}[(\bar{X} - \bar{\mu})(\bar{X} - \bar{\mu})^T] \bar{a}$$

$$= \mathbb{E} \left[\underbrace{\bar{a}^T (\bar{X} - \bar{\mu})}_{z} \underbrace{(\bar{X} - \bar{\mu})^T \bar{a}}_{z^T = z} \right] = \mathbb{E}[z^2] \geq 0$$

Application: If $K_{\bar{X}}$ is positive semi-definite, then it has a "square-root", i.e. $K_{\bar{X}} = SS^T$ if $a \geq 0$, then $a = r^2$ for $r \in \mathbb{R}$.

$$\Lambda^{1/2} = \begin{bmatrix} \sqrt{\lambda_1} & & 0 \\ & \sqrt{\lambda_2} & \\ 0 & & \ddots \\ & & & \sqrt{\lambda_n} \end{bmatrix}$$

$$K_{\bar{X}} = \underbrace{Q \Lambda^{1/2}}_S \underbrace{\Lambda^{1/2} Q^T}_{S^T}$$

$$S^T = (Q \Lambda^{1/2})^T = (\Lambda^{1/2})^T Q^T$$

Coloring:

\bar{X} with $K_{\bar{X}} = I$, i.e. white.

We want to create \bar{Y} with covariance $K_{\bar{Y}}$

$$K_{\bar{Y}} = SS^T$$

$$\bar{Y} = S\bar{X} \quad K_{\bar{Y}} = SS^T$$

Whitening:

\bar{X} with covariance $K_{\bar{X}} = SS^T$

$$\bar{Y} = S^{-1/2} \bar{X}$$

$$K_{\bar{Y}} = S^{-1/2} \underbrace{SS^T (S^{-1/2})^T}_{(S^{-1/2} S)^T} = I$$

assuming S is invertible